

Statistical functionals and influence functions

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Recap

In our first lecture, we introduced the empirical distribution function as a nonparametric way of estimating F , and showed that it had a number of very attractive properties:

- $\hat{F} \xrightarrow{\text{a.s.}} F$, regardless of F
- The nonparametric maximum likelihood estimator
- Able to derive confidence intervals and confidence bands that work for any F , any n , and all x

Parametric vs. nonparametric estimation

- Why so much emphasis on estimating F ?
- Because \hat{F} will play the same role in nonparametric estimation that $\hat{\theta}$ played in parametric estimation
- In parametric statistics, we find the most likely distribution of the data based on $\hat{F} = F(\hat{\theta})$
- This then provides estimates of our parameters of interest – usually θ itself or some function of θ , such as e^{θ} , but conceivably also something like the 75th percentile

Nonparametric estimation

- Clearly, nonparametric statistics does not provide estimates of parameters (or functions of parameters) . . . so what are we estimating in nonparametric statistics?
- We estimate what are called *statistical functionals*: a functional $\theta = T(F)$ is any function of F
- Many common descriptive statistics can be expressed as statistical functionals

Examples

- Mean: $T(F) = \int x dF(x)$
- Variance: $T(F) = \int (x - \mu)^2 dF(x)$
- Quantiles: $T(F) = F^{-1}(p)$

Estimation of statistical functionals

- Estimation of these quantities is straightforward: calculate the quantity you are interested in based on the nonparametric maximum likelihood estimator of F , the empirical distribution function \hat{F} :

$$\hat{\theta} = T(\hat{F})$$

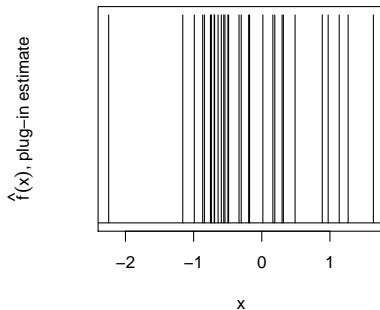
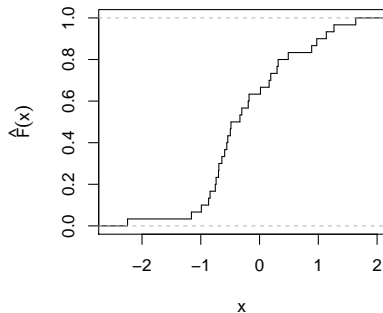
- This idea is often called the *plug-in principle*, and its resulting estimate the *plug-in estimate*
- Examples:
 - Mean: \bar{x}
 - Variance: $n^{-1} \sum (x_i - \bar{x})^2$
 - Quantile: Sample quantile
- In a sense, the plug-in principle is a nonparametric analog of the likelihood principle

Is the plug-in estimator a good estimator?

- A natural question: Is the plug-in estimator a good estimator?
- The Glivenko-Cantelli Theorem says that $\hat{F} \xrightarrow{\text{a.s.}} F$; does this mean that $T(\hat{F}) \xrightarrow{\text{a.s.}} T(F)$?
- The answer turns out to be a complicated “sometimes”; often “yes”, but not always

Plug-in density estimation

As an example of a case where the plug-in estimator is not consistent, consider density estimation, where $T(F) = F'(x)$:



The influence function

- So when is the plug-in estimator consistent?
- It requires certain conditions on the smoothness (differentiability) of $T(F)$
- But what does it mean to take the derivative of a function with respect to a function?
- To answer this question, we will have to expand our notion of the derivative

The Gâteaux derivative

- The extension of the derivative that we will need is called the *Gâteaux derivative*
- The Gâteaux derivative of T at F in the direction G is defined by

$$L_F(T; G) = \lim_{\epsilon \rightarrow 0} \left[\frac{T\{(1 - \epsilon)F + \epsilon G\} - T(F)}{\epsilon} \right]$$

- An equivalent way of stating the definition is to define $D = G - F$, and the above becomes

$$L_F(T; D) = \lim_{\epsilon \rightarrow 0} \left[\frac{T\{F + \epsilon D\} - T(F)}{\epsilon} \right]$$

- Either way, the definition boils down to

$$L_F(T) = \lim_{\epsilon \rightarrow 0} \left[\frac{T(F_\epsilon) - T(F)}{\epsilon} \right]$$

Interpretation

- From a mathematical perspective, the Gâteaux derivative a generalization of the concept of a directional derivative to functional analysis
- From a statistical perspective, it represents the rate of change in a statistical functional upon a small amount of contamination by another distribution G

Gâteaux derivative of $f(x)$

- As an example of how the Gâteaux derivative works, suppose F is a continuous CDF, and G is the distribution that places all of its mass at the point x_0
- What happens to the Gâteaux derivative of $T(F) = f(x_0)$?

$$\begin{aligned} L_F(T; G) &= \lim_{\epsilon \rightarrow 0} \left[\frac{\frac{d}{dx} \{(1 - \epsilon)F(x) + \epsilon G(x)\}_{x=x_0} - \frac{d}{dx} F(x)|_{x=x_0}}{\epsilon} \right] \\ &= \lim_{\epsilon \rightarrow 0} \left[\frac{(1 - \epsilon)f(x_0) + \epsilon g(x_0) - f(x_0)}{\epsilon} \right] \\ &= \infty \end{aligned}$$

Glivenko-Cantelli does not imply convergence of estimators

- So, even though F and F_ϵ differ from each other only infinitesimally $T(F)$ and $T(F_\epsilon)$ differ from each other by an infinite amount
- Thus, the Glivenko-Cantelli theorem does not help us here:
 $\sup_x |\hat{F}(x) - F(x)|$ may go to zero without $T(\hat{F}) \rightarrow T(F)$

Hadamard differentiability

- It turns out that even Gâteaux differentiability is too weak to ensure that $T(\hat{F}) \rightarrow T(F)$
- Even if the Gâteaux derivative exists, it may not exist in an entirely unique way, and this is the subtle idea introduced by *Hadamard differentiability*
- A functional T is Hadamard differentiable if, for any sequence $\epsilon_n \rightarrow 0$ and D_n satisfying $\sup_x |D_n(x) - D(x)| \rightarrow 0$, we have

$$\frac{T(F + \epsilon_n D_n) - T(F)}{\epsilon_n} \rightarrow L_F(T; D)$$

- If T is Hadamard differentiable, then $T(\hat{F}) \xrightarrow{P} T(F)$

Bounded functionals

Another useful condition that — certainly, one that is easier to check than Hadamard differentiability — is that if the functional is bounded, then the plug-in estimate will converge to the true value

Homework: Suppose that there exists a constant C such that the following relation holds for all G :

$$|T(F) - T(G)| \leq C \sup_x |F(x) - G(x)|.$$

Show that $T(\hat{F}) \xrightarrow{\text{a.s.}} T(F)$.

Contamination by a point mass

- The idea of contaminating a distribution with a small amount of additional data has a long history in statistics and the investigation of robust estimators
- Statisticians usually do not work with the general Gâteaux derivative, but a special case of it called the *influence function*, in which G places a point mass of 1 at x :

$$\delta_x(u) = \begin{cases} 0 & \text{if } u < x \\ 1 & \text{if } u \geq x \end{cases}$$

The influence function and empirical influence function

- The influence function is usually written as a function of x , and defined as

$$L(x) = \lim_{\epsilon \rightarrow 0} \left[\frac{T\{(1 - \epsilon)F + \epsilon\delta_x\} - T(F)}{\epsilon} \right]$$

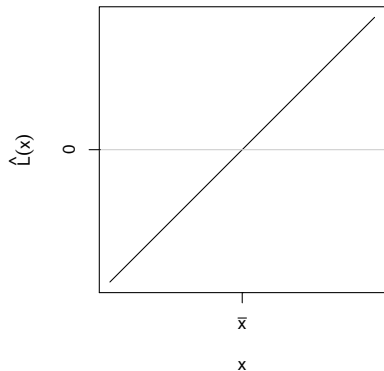
- A closely related concept is that of the *empirical influence function*:

$$\hat{L}(x) = \lim_{\epsilon \rightarrow 0} \left[\frac{T\{(1 - \epsilon)\hat{F} + \epsilon\delta_x\} - T(\hat{F})}{\epsilon} \right]$$

Example: The mean

$$L(x) = x - \mu$$

$$\hat{L}(x) = x - \bar{x}$$



Linear functionals

- The mean is an example of a *linear functional*: one in which

$$T(F) = \int a(x) dF(x)$$

- Linear functionals are particularly easy to work with, as

$$L(x) = a(x) - T(F)$$

$$\hat{L}(x) = a(x) - T(\hat{F})$$

The chain rule

- Working directly from the definition each time is time-consuming
- Fortunately, the Gâteaux derivative has many of the same properties as ordinary derivatives – in particular, the chain rule
- Suppose our functional can be written in the form $T(F) = a\{T_1(F), T_2(F), \dots\}$; then

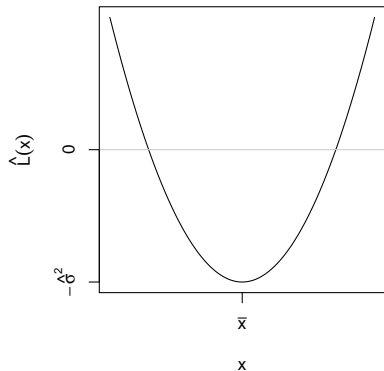
$$L(x) = \sum_j \left. \frac{\partial a}{\partial T_j} \right|_F L_j(x),$$

where $L_j(x)$ is the influence function of $T_j(F)$

Example: Variance

$$L(x) = (x - \mu)^2 - \sigma^2$$

$$\hat{L}(x) = (x - \bar{x})^2 - \hat{\sigma}^2$$



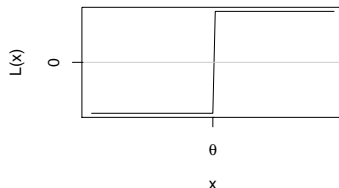
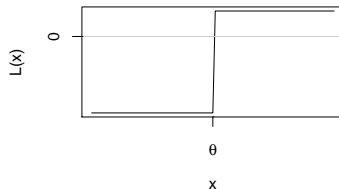
Homework

Homework: Consider a random variable X that is always positive. We are interested in the statistical functionals $\theta = \int \log(x) dF(x)$ and $\lambda = \log(\mu)$.

- (b) What are the influence and empirical influence functions for θ ?
- (d) What are the influence and empirical influence functions for λ ?
- (f) Do $\hat{\lambda}$ and $\hat{\theta}$ converge to the same number?
- (g) Plot the empirical influence functions from parts (b) and (d). Label the point x on the horizontal axis where $L(x) = 0$.
- (h) Briefly, comment on the relative robustness of $\hat{\theta}$ and $\hat{\lambda}$ to outliers.

Example: Quantiles

$$L(x) = \begin{cases} \frac{p-1}{f(\theta)} & \text{if } x \leq \theta \\ \frac{p}{f(\theta)} & \text{if } x > \theta \end{cases}$$
$$\hat{L}(x) = ?$$

Median**25th percentile**

Homework

Homework: Show that, for $x > \theta$,

$$L(x) = \frac{p}{f(\theta)}$$

Breakdown point

- Note that the influence function for the median is unbounded, while the influence function for the mean is not
- **Homework:** Let $b(\epsilon) = \sup_x |T(F) - T(F_\epsilon)|$, where $F_\epsilon = (1 - \epsilon)F + \epsilon\delta_x$. The *breakdown point* of an estimator, ϵ^* , is defined as $\epsilon^* = \inf\{\epsilon : b(\epsilon) = \infty\}$
 - a) Find the breakdown point of the mean
 - b) Find the breakdown point of the median