High-Dimensional Inference for Sparse Vector Autoregressions (Sparse VAR)

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The data, $\{\mathbf{x}_1, \ldots, \mathbf{x}_n\}$, is postulated to come from an appropriate model class, $\mathcal{M} = \{f(\mathbf{x}|\boldsymbol{\theta}) : \boldsymbol{\theta} \in \Theta \subset \mathbb{R}^d\}$.

- Model Selection*: apply sound principles to select an appropriate element from \mathcal{M} ; must resolve the bias-variance tradeoff problem...
 - use an information criterion (AIC/BIC); etc.
- Model Fitting: straightforward minimization of an appropriate criterion to estimate $\theta \mapsto \hat{\theta}$:
 - least-squares (LSE); maximum likelihood (MLE); etc.
- Model Checking: residuals from the fitted ("true") model have predictable properties:
 - independence; homoscedasticity; (asymptotically) normal; etc.
- Inference*: Confidence intervals & hypothesis tests on functions of θ; usually based on an asymptotic normality result:

$$\sqrt{n}(\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}) \stackrel{d}{\rightarrow} N(\boldsymbol{0}, \Omega)$$

- VAR Models: Ubiquitous in (stationary) time series modeling.
- Goal: Model MANY series simultaneously (high dimensional)

VAR dimension = d

- Existing Methods: Lasso; Elasticnet; SCAD; MCP; shrinkage priors (Bayesian)
- R Packages: sparsevar; BigVAR
- New Methods: p-value thresholding; BIC-based forward selection
- Simulations: Compare sparsity pattern recovery and accuracy
- Illustrations: fMRI data, S&P 500 Data



Macroeconomic Forecasting (Kastner & Huber, 2019): d = 215 quantities



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$$\mathbf{x}_t = \Phi_1 \mathbf{x}_{t-1} + \cdots + \Phi_p \mathbf{x}_{t-p} + \boldsymbol{\epsilon}_t, \qquad \{\boldsymbol{\epsilon}_t\} \sim \mathsf{iid}(\mathbf{0}, \boldsymbol{\Sigma}_{\epsilon})$$

stationary/stable/causal

- G is Toeplitz matrix of first p process autocovariances
- total number of parameters/coefficients: $m := pd^2$

$$\psi = \mathsf{vec}([\Phi_1, \dots, \Phi_p])$$

• Index set of all elements of ψ :

$$I = \{1, \ldots, m\} = J \cup K$$

where

•
$$J =$$
active elements ($\neq 0$)

• *K* = non-active elements (= 0)

Both least-squares (LSEs) & max likelihood (MLEs) have same asymptotics:

$$\sqrt{n}(\hat{\psi} - \psi) \stackrel{d}{
ightarrow} (\mathbf{0}, \Omega), \qquad \Omega = \Sigma_{\epsilon} \otimes G^{-1}$$

- \hat{p}_i : p-value for testing null that $\psi_i = 0$
- Ordered p-values:

$$\hat{p}_{(1)} \leq \cdots \leq \hat{p}_{(m)}$$

- A1. The VAR is strictly stationary and stable, and the covariance matrix $\Sigma_{\epsilon} \otimes G^{-1}$ is well defined and non-singular.
- A2. The iid series $\{\epsilon_t = (\epsilon_{1t}, \ldots, \epsilon_{dt})'\}$ has a continuous distribution with a positive definite covariance Σ_{ϵ} , and satisfies $\mathbb{E}|\epsilon_{it}\epsilon_{jt}\epsilon_{kt}\epsilon_{lt}| < \infty$ for any $i, j, k, l \in \{1, \ldots, d\}$, and all t.
- A3. The total number of coefficient parameters m is such that m < n for all sufficiently large n, and the threshold activation parameter is of order: $h_n = O(\log n)$.
- A4. The threshold satisfies $p_n \to 0$ and $\log(p_n) = o(n)$ as $n \to \infty$.

Declare as **non-active** coefficients for which p-value exceeds threshold:

 $\hat{p}_i > p_n$

Theorem

Let $\psi_{J_0} \neq \mathbf{0}$ denote the vector of active parameters of the stationary and stable VAR model, and $\hat{\psi}_{\hat{J}}$ its TLSE estimator. Then, under conditions **A1–A4**, we have as $n \to \infty$:

•
$$P(\hat{J}=J_0) \rightarrow 1$$

•
$$\sqrt{n}(\hat{\psi}_{\hat{J}} - \psi_{J_0}) \stackrel{d}{\rightarrow} N(\mathbf{0}, \Omega_{J_0})$$

Declare as **non-active** coefficients for which p-value exceeds threshold:

 $\hat{p}_i > p_n$

Consider following t-holds (satisfy A3 & A4):

• TLSE1:
$$p_n = \frac{1}{\sqrt{h_n}\log(h_n)}$$

• TLSE2: $p_n = \frac{1}{\sqrt{h_n\log(h_n)}}$
• TLSE3: $p_n = \frac{h_n}{n\log(h_n)}$

Activation parameter choice:

$$h_n = m \log n$$

• Ordered p-values and corresponding concomitant VAR coefficients:

$$\{\hat{p}_{(1)} \leq \cdots \leq \hat{p}_{(m)}\} \quad \iff \quad \{\psi_{(1)}, \dots, \psi_{(m)}\}$$

• For $j = 1, \ldots, m$: model $\widehat{\mathcal{M}}_j$ contains coefficients

$$\{\psi_{(1)},\ldots,\psi_{(j)}\}$$

• Nested sequence of models:

$$\widehat{\mathcal{M}}_1 \subset \cdots \subset \widehat{\mathcal{M}}_m$$

• BLSE: Model with smallest BIC in nested sequence

 $\widehat{\mathcal{M}}_*$

BLSE Algorithm:

- delivers Model $\widehat{\mathcal{M}}_*$
- identifies the (active) index set $\widehat{J_*}$

Theorem

For the stationary and stable VAR(p) with Assumptions A1-A4:

$$P(J_0\subset \widehat{J_*}) o 1, \qquad \text{as } n o \infty$$

- Models (2):
 - Medium VAR: d = 10 dim VAR(1), n = 200.
 - Large VAR: d = 30 dim VAR(1), n = 2000.
- Sparsity Values (3): $s = \{0.2, 0.5, 0.8\}$.
- Methods (9):
 - L1-LS: OLS-based lasso (sparsevar)
 - L1-LL: MLE-based lasso (sparsevar)
 - BVEN: OLS-based elastic net (BigVAR)
 - BVSCAD: OLS-smoothly clipped absolute deviation (BigVAR)
 - BVMCP: OLS-minimax concave penalty (BigVAR)
 - BLSE
 - TLSE1
 - TLSE2
 - TLSE3

- **Sparsity pattern recovery.** Use usual metrics from binary classification:
 - *acc*₁: true positive rate (sensitivity)
 - acco: true negative rate (specificity)
 - acc: overall accuracy (proportion correctly classified)

• Coefficient estimation.

- RMSE: only for the truly active coefficients
- 1-Step forecast error.
 - RMSFE

(All based on 100 replications)











time

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Table: Percentage of non-active coefficients (sparsity) for the Φ_1 matrix of VAR(1) models fitted by each method.

TLSE1	TLSE2	TLSE3	BLSE	BVEN	BVMCP	BVSCAD	L1-LL	L1-LS
90	86	54	36	27	7	7	16	45

Brain fMRI Data: Sparsity Pattern Distances (Scaled Manhattan Metric)





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TLSE1	TLSE2	TLSE3	BLSE	BVEN	BVMCP	BVSCAD	L1-LL	L1-LS
77	74	69	99	34	100	100	100	51

S&P 500 Data: Sparsity Pattern Distances (Scaled Manhattan Metric)



Granger (1969):

Y Granger-Causes X if prediction of X improves when one uses past values of Y, given that all other relevant information, Z, is taken into account.

For VAR Model (Geweke, 1984):

- $\Sigma_{XY|Z}$: prediction error covariance for X|Z when Y is included
- $\Sigma_{XX|Z}$: prediction error covariance for X|Z when Y is excluded

$$GC(Y \to X|Z) = \log\left(\frac{|\Sigma_{XX|Z}|}{|\Sigma_{XY|Z}|}
ight).$$

GC(Y → X|Z) > 0: suggests Y is G-Causal for X
GC(Y → X|Z) ≤ 0: suggests Y is not G-Causal for X

Brain fMRI Data: GC-Flow



GC Strength





BLSE



Value

0.4

0.2

0.0 -0.2

-0.4



S&P 500 Data: GC-Flow

GC Strength



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Thank You!