### 2 Discrete-Time Markov Chains

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Biomathematics II MATH 5355 Spring 2017

Lecture notes follow: Allen, Linda JS. An introduction to stochastic processes with applications to biology. CRC Press, 2010.

Discrete-time stochastic process  $\{X_n\}_{n=0}^{\infty}$ 

- X<sub>n</sub> is discrete random variables on finite or countably infinite state space
- *n* index is used for time  $\{0, 1, 2, ...\}$

### Discrete-Time Markov Chain (DTMC)

 $\{X_n\}_{n=0}^\infty$  has the Markov property if

$$\mathsf{Prob}\{X_n = i_n | X_0 = i_0, ..., X_{n-1} = i_{n-1}\} = \mathsf{Prob}\{X_n = i_n | X_{n-1} = i_{n-1}\}$$

and the process is called a DTMC.

- ► Notation: Prob{·} = P<sub>X<sub>n</sub></sub>{·} used. (P will represent a transition matrix).
- ►  $\{p_i(n)\}_{n=0}^{\infty}$  is the p.m.f. associate with  $X_n$ , where  $p_i(n) = \text{Prob}\{X_n = i\}.$
- ► Transition probabilities relate state of process at time n to n+1 (X<sub>n</sub> to X<sub>n+1</sub>).

One-step Transition Probability

$$p_{ji}(n) = \operatorname{Prob}\{X_{n+1} = j | X_n = i\}$$

is the probability that the process is in state j at time n + 1 given that the process was in state i at time n. For each state,  $p_{ii}$  satisfies

$$\sum_{j=1}^\infty p_{ji} = 1$$
 &  $p_{ji} \ge 0.$ 

- The above summation means the process at state i must transfer to j or stay in i during the next time interval.
- *p<sub>ji</sub>* don't depend on time they are *stationary* or *homogeneous*
- ▶  $p_{ji}(n)$  do depend on time they are *nonstationary* or *nonhomogeneous*

### Transition Matrix

The DTMC  $\{X_n\}_{n=0}^{\infty}$  with one-step transition probabilities  $\{p_{ij}\}_{i,j=1}^{\infty}$  has transition matrix  $P = (p_{ij})$ :

	$(p_{11})$	$p_{12}$		$p_{1j}$	)
P =	<i>p</i> <sub>21</sub>	<i>p</i> <sub>22</sub>		$p_{2j}$	
	÷	÷	·	÷	·
	<i>p</i> <sub>i1</sub>	$p_{i2}$		p <sub>ij</sub>	
	( :	÷	۰.	÷	·.)

- Columns sum to 1, since  $\sum_{j=1}^{\infty} p_{ji} = 1$ .
- Called a *Stochastic Matrix*.
- Note notation: p<sub>ij</sub> is the probability of transition from state j to state i (other sources may define this differently).

N-step Transition Probability

$$p_{ji}^{(n)} = \operatorname{Prob}\{X_n = j | X_0 = i\}$$

is the probability of transferring from state *i* to state *j* in *n* time steps. The n-step transition matrix  $P^{(n)} = \left(p_{ji}^{(n)}\right)$ , where  $p_{ji}^{(1)} = p_{ji}$  and

$$p_{ji}^{(0)} = \delta_{ji} = \begin{cases} 1, & j = i \\ 0, & j \neq i \end{cases}$$

where  $\delta_{ji}$  is the Kronecker delta symbol. Then  $P^{(1)} = P$  and  $P^{(0)} = I$  the identity matrix.

Chapman-Kolmogorov Equations

$$p_{ji}^{(n)} = \sum_{k=1}^{\infty} p_{jk}^{(n-s)} p_{ki}^{(s)}, \qquad 0 < s < n$$

Or in terms of matrix notation:

$$P^{(n)} = P^{(n-s)}P^{(s)}$$

Here,

$$P^{(1)} = P$$
  
 $P^{(2)} = P^{(1)}P^{(1)} = P^2$   
:  
 $P^{(n)} = P^n$ 

The n-step transition matrix  $P^{(n)}$  is just the nth power of P.

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Let  $p(n) = (p_1(n), p_2(n), ...)^T$  be the vector form of the p.m.f. for  $X_n$  where  $p_i(n) = \text{Prob}\{X_n = i\}$ . The probabilities satisfy

$$\sum_{i=1}^{\infty} p_i(n) = 1.$$

The probability distribution associated with  $X_{n+1}$  can be found:

$$p_i(n+1) = \sum_{j=1}^{\infty} p_{ij}p_j(n)$$
 or  $p(n+1) = Pp(n)$ 

This projects the process *forward* in time.

$$p(n+m) = P^{n+m}p(0) = P^n(P^mp(0)) = P^np(m)$$

# Example

Etterson et al. 2009<sup>\*</sup> propose a simple Markov chain model to describe the reproductive activities of a single female bird in a single breeding season. A female can occupy one of four states:

- 1. actively nesting
- 2. successfully fledged a brood
- 3. failed to fledge a brood
- 4. completed all nesting activities for the season



The state space is  $E = \{1, 2, 3, 4\}$  and the random variable  $X_n \in E$  represents the state of the female following the nth change of state.

<sup>\*</sup>Etterson, Matthew A., et al. "Markov chain estimation of avian seasonal fecundity." Ecological Applications 19.3 (2009): 622-630.

## Example

Etterson et al. 2009 used the transition matrix:

$$P = \begin{pmatrix} 0 & s^a & 1 - s^a & 0 \\ 1 - q_s & 0 & 0 & q_s \\ 1 - q_f & 0 & 0 & q_f \\ 0 & 0 & 0 & 1 \end{pmatrix} = \begin{pmatrix} 0 & 0.369 & 0.631 & 0 \\ 0.33 & 0 & 0 & 0.67 \\ 0.58 & 0 & 0 & 0.42 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

where *s* is the daily nest survival probability, *a* is the average time from first egg to fledging,  $q_s$  is the probability that a female quits breeding following a successful breeding attempt, and  $q_f$  is the probability that a female quits breeding following a failed breeding attempt. The numbers are estimates obtained from field studies of a population of Eastern Meadowlarks in Illinois.

### Communicate

State *j* can be reached from state *i* is there is a nonzero probability  $p_{ji}^{(n)} > 0$  for some  $n \ge 0$ . This is denoted as  $i \to j$ . If  $i \to j$  and  $j \to i$ , *i* and *j* are said to communicate, or be in the same class, denoted as  $i \leftrightarrow j$ ; that is, there exists *n* and *n'* such that

$$p_{ji}^{(n)} \ 0 \quad \& \quad p_{ij}^{(n')} > 0$$

Directed Graph:



Here,  $i \to j$  as  $p_{ji} > 0$  and  $i \to k$  as  $p_{ki}^{(2)} > 0$  but is is not that case that  $k \to i$ .

### Equivalence Relation

 $i \leftrightarrow j$  is an equivalence relation on the state  $\{1,2,...\}$ 

- 1. reflexivity:  $i \leftrightarrow i$  (because  $p_{ii}^{(0)} = 1$ )
- 2. symmetry:  $i \leftrightarrow j$  implies  $j \leftrightarrow i$
- 3. transitivity:  $i \leftrightarrow j, j \leftrightarrow k$  implies  $i \leftrightarrow k$ .

### Communication Classes

The set of equivalences classes in a DTMC are the communication classes. If every state in the Markov chain can be reached by every other state, then there is only one communication class.

#### Irreducible

If there is only one communication class, then the Markov chain is irreducible, otherwise is it reducible.

### Irreducible

A transition matrix P is irreducible if the directed graph is strongly connected. It is reducible if the directed graph is not strongly connected.

#### Closed

Set of states C is closed if it is impossible to reach any state outside of C from any state in C by one-step transitions:  $p_{ji} = 0$  if  $i \in C$  and  $j \notin C$ .

If C is a closed communicating class for a Markov chain X, then that means that once X enters C, it never leaves C.

#### Absorbing State

State *i* is absorbing if  $p_{ii} = 1$ .

If i is an absorbing state once the process enters state i, it is trapped there forever.



The Markov chain is irreducible and it periodic with period N (beginning in state *i*, it takes N steps to return to state *i*:  $P^N = I$ )

Periodic

The period of state *i* is the greatest common divisor of all  $n \ge 1$  for which  $p_{ii}^{(n)} > 0$ :

$$d(i) = g.c.d.\{n|p_{ii}^{(n)} > 0 \text{ and } n \ge 1\}.$$

If d(i) > 1 the state is periodic of period d(i). If d(i) = 1 the state is aperiodic. If  $p_{ii}^{(n)} = 0$  for all  $n \ge 1$  define d(i) = 0.

#### Example

 $P = \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix}$ 

Communication classes:  $\{1\}, \{2\}, \{3\}$ . d(i) = 0 for i = 1, 2 because here  $p_{ii}^{(n)} = 0$ . State 3 is aperiodic since d(3) = 1.

### Example



3 communication classes:  $\{1\}, \{3\}, \{2,4\}$ . Markov chain is reducible.

#### Etterson et al. 2009

Etterson et al. 2009 used the transition matrix:

$$P = egin{pmatrix} 0 & s^a & 1-s^a & 0 \ 1-q_s & 0 & 0 & q_s \ 1-q_f & 0 & 0 & q_f \ 0 & 0 & 0 & 1 \end{pmatrix}$$

where  $s \in (0, 1)$ , a > 0, and  $q_s$ ,  $q_f \in (0, 1)$ . State space  $E = \{1, 2, 3, 4\}$  includes two communicating classes:  $C_1 = \{1, 2, 3\}$  and  $C_2 = \{4\}$ . The Markov chain is reducible.  $C_2$  is closed and 4 is an absorbing state.

#### First Return Probability

Let  $f_{ii}^{(n)}$  be the probability that starting from state *i*,  $X_0 = i$ , the first return to state *i* is at the nth time step:

$$f_{ii}^{(n)} = \text{Prob}\{X_n = i, X_m \neq i, m = 1, 2, ..., n - 1 | X_0 = i\}, n \ge 1.$$

The probabilities  $f_{ii}^{(n)}$  are the **first return probabilities.**  $f_{ii}^{(0)} = 0$ .  $f_{ii}^{(1)} = p_{ii}$ , but generally  $f_{ii}^{(n)} \neq p_{ii}^{n}$ .

Recurrent State State *i* is recurrent if

$$\sum_{n=1}^{\infty} f_{ii}^{(n)} = 1.$$

Let random variable  $T_{ii}$  be the first return time. Then  $\{f_{ii}^{(n)}\}_{n=0}^{\infty}$  defines a probability distributions for  $T_{ii}$ . Here,  $T_{ii} = n$  with probability  $f_{ii}^{(n)}$ .

Transient State

State *i* is transient if

$$\sum_{n=1}^{\infty} f_{ii}^{(n)} < 1.$$

Then  $\{f_{ii}^{(n)}\}_{n=0}^{\infty}$  is not a complete set of probabilities needed to define a probability distribution. Here, let  $f_{ii} = \sum_{n=0}^{\infty} f_{ii}^{(n)} < 1$  and define  $1 - f_{ii}$  as the probability of never returning to *i*.  $T_{ii}$  is the *waiting time* until the chain returns to *i*.

### Mean Recurrence Time

For recurrent state *i*, the mean recurrence time is the mean of the distribution  $T_{ii}$ :

$$u_{ii} = E(T_{ii}) = \sum_{n=1}^{\infty} n f_{ii}^{(n)}$$

For the recurrent state *i* if  $\mu_{ii} < \infty$  it is positive recurrent. If  $\mu_{ii} = \infty$ , it is null recurrent.

- Ex: absorbing state *i*. Here,  $p_{ii} = 1$  thus  $f_{ii}^{(1)} = 1$  and  $f_{ii}^{(n)} = 0$  for  $n \neq 1$ . Therefore  $\mu_{ii} = 1$ .
- ▶ The mean recurrence time for a transient state is infinity:  $T_{ii} = \infty$  with probability  $1 f_{ii}$

### Example

Consider a two state Markov chain with transition matrix:

$$P = egin{pmatrix} p_{11} & p_{12} \ p_{21} & p_{22} \end{pmatrix}$$

where  $0 < p_{ii} < 1$  for i = 1, 2. Show that both states are positive recurrent.

# First Passage Time

### First Passage Time Probability

Let  $f_{ji}^{(n)}$  be the probability that starting from state *i*,  $X_0 = i$ , the first return to state *j* is at the nth time step:

$$f_{ji}^{(n)} = \text{Prob}\{X_n = j, X_m \neq j, m = 1, 2, ..., n - 1 | X_0 = i\}, \quad j \neq i, n \ge 1.$$

The probabilities  $f_{ji}^{(n)}$  are the **first passage time probabilities.**  $f_{ji}^{(0)} = 0$ .

### First Passage Form State *j* from State *i*

If  $\sum_{n=0}^{\infty} f_{ji}^{(n)} = 1$  then  $\{f_{ji}^{(n)}\}$  defines a probability distribution for a random variable  $T_{ji}$ , the first passage to state *j* from state *i*.

### Mean First Passage Time

If  $X_0 = i$ , the mean first passage time to state j is:  $\mu_{ji} = E(T_{ji}) = \sum_{n=1}^{\infty} nf_{ji}^{(n)}, \ j \neq i$  Relationships between step transition and first return probabilities:

$$p_{ii}^{(n)} = \sum_{k=1}^{n} f_{ii}^{(k)} p_{ii}^{(n-k)}$$
 &  $p_{ji}^{(n)} = \sum_{k=1}^{n} f_{ji}^{(k)} p_{ji}^{(n-k)}$ 

#### Generating Functions

Let the generating function for the sequence  $\{f_{ii}^{(n)}\}$  be

$$F_{ji}(s) = \sum_{n=0}^{\infty} f_{ji}^{(n)} s^n, \quad |s| < 1$$

Let the generating function for the sequence  $\{p_{ji}^{(n)}\}$  be

$$extsf{P}_{ji}(s) = \sum_{n=0}^\infty p_{ji}^{(n)} s^n, \quad |s| < 1$$

Relationships:  $F_{ii}(s)P_{ii}(s) = P_{ii}(s) - 1$  &  $F_{ji}(s)P_{jj}(s) = P_{ji}(s)$ 

### Basic Theorems for Markov Chains

Theorem 2.2

A state *i* is recurrent (transient) if and only if  $\sum_{n=0}^{\infty} p_{ii}^{(n)}$  diverges (converges), i.e.

$$\sum_{n=0}^{\infty} p_{ii}^{(n)} = \infty (<\infty)$$

The Proof of this used the following theorem:

Abel's Convergence Theorem

If 
$$\sum_{k=0}^{\infty} a_k$$
 converges, then  $\lim_{s \to 1^-} \sum_{k=0}^{\infty} a_k s^k = \sum_{k=0}^{\infty} a_k = a$   
If  $a_k \ge 0$  and  $\lim_{s \to 1^-} \sum_{k=0}^{\infty} a_k s^k = a \le \infty$ , then  $\sum_{k=0}^{\infty} a_k = a$ 

# Basic Theorems for Markov Chains

### Theorem 2.2

A state *i* is recurrent (transient) if and only if  $\sum_{n=0}^{\infty} p_{ii}^{(n)}$  diverges (converges), i.e.

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### Corollaries

- ► Assume i ↔ j. State i is recurrent (transient) if and only if state j is recurrent (transient)
- Every recurrent class in a DTMC is a closed set

# Infinite Random Walk



p > 0 is the probability of moving to the right:  $p_{i+1,i} = p$ q > 0 is the probability of moving to the left:  $p_{i,i+1} = q$ p + q = 1

### **Example Questions**

- 1. Is the MC reducible or irreducible?
- 2. Is the MC aperiodic or periodic?
- 3. Assume  $p = q = \frac{1}{2}$ , is the MC transient or recurrent?
- 4. Assume  $p \neq q$ , is the MC transient or recurrent?

### Basic Theorems for Markov Chains

### Basic Limit Theorem, Aperiodic Markov Chains

Let  $\{X_n\}_{n=0}^{\infty}$  be a recurrent, irreducible, and aperiodic DTMC with transition matrix  $P = (p_{ij})$ :

$$\lim_{n\to\infty}p_{ij}^{(n)}=\frac{1}{\mu_{ii}}$$

#### Basic Limit Theorem, Periodic Markov Chains

Let  $\{X_n\}_{n=0}^{\infty}$  be a recurrent, irreducible, and d-periodic DTMC, d > 1, with transition matrix  $P = (p_{ij})$ :

$$\lim_{n\to\infty}p_{ii}^{(nd)}=\frac{d}{\mu_{ii}}$$

Summary of Classification Schemes

Markov chains or classes can be classified as

# Periodic or Aperiodic

Then further classified as

## Transient or Recurrent

Then recurrent MC can be classified as

Null recurrent or Positive recurrent.

"Equilibirum" of the Markov Chain

### Stationary Probability Distribution

A stationary probability distribution is the vector  $\pi = (\pi_1, \pi_2, ...)^T$  with:

$$P\pi = \pi$$
 &  $\sum_{i=1}^{\infty} \pi_i = 1.$ 

For a finite MC  $\pi$  is an eigenvector of P with eigenvalue  $\lambda = 1$ :

$$P\pi = \lambda \pi$$
 &  $\sum_{i=1}^{N} \pi_i = 1$ 

If a chain is initially at a stationary probability distribution  $p(0) = \pi$ , then  $p(n) = P^n \pi = \pi$  for all time *n*.

# "Equilibirum" of the Markov Chain

There may be more than one linearly independent eigenvector for  $\lambda = 1$ . In this case, the stationary probability distribution is not unique. However a positive recurrent, irreducible, and aperiodic DTMC has a unique stationary probability distribution:

### Theorem 2.5

Let  $\{X_n\}_{n=0}^{\infty}$  be a positive recurrent, irreducible, and aperiodic DTMC. There is a unique positive stationary probability distribution  $\pi$  with  $P\pi = \pi$ 

$$\lim_{n \to \infty} p_{ij}^{(n)} = \pi_i, \qquad i, j = 1, 2, \dots$$

The basic limit theorem then yields:

$$\pi_i=\frac{1}{\mu_{ii}}>0$$

where  $\mu_{ii}$  is the mean recurrence time for state *i*.

# "Equilibirum" of the Markov Chain

### Example

$$P=egin{pmatrix} 1/2 & 1/3 \ 1/2 & 2/3 \end{pmatrix}$$

- 1. What is the stationary probability distribution for P?
- 2. What are the mean recurrence times?

In finite DTMC, there are **NO null recurrent states** and **not all states** can be transient.

- 4 classification schemes: periodic or aperiodic and transient or positive recurrent
- An irreducible finite DTMC is positive recurrent
- In a finite DTMC, a class is recurrent if and only if it is closed

# Grey vs. Red Squirrel

Red Squirrel are native to areas of Great Britain and Gray Squirrels invaded many regions in the  $19^{th}$  century. Each region is classified as being in one the following states:

- 1. occupied by Red squirrels only
- 2. occupied by Gray squirrels only
- 3. occupied by both
- 4. no squirrels

The transitions between states over a period of 1 year were estimated for the following transition matrix:

$$P = \begin{pmatrix} 0.8797 & 0.0382 & 0.0527 & 0.008\\ 0.0212 & 0.8002 & 0.0041 & 0.0143\\ 0.0981 & 0.0273 & 0.8802 & 0.0527\\ 0.0010 & 0.1343 & 0.0630 & 0.9322 \end{pmatrix}$$

### Grey vs. Red Squirrel

$$P = \begin{pmatrix} 0.8797 & 0.0382 & 0.0527 & 0.0008 \\ 0.0212 & 0.8002 & 0.0041 & 0.0143 \\ 0.0981 & 0.0273 & 0.8802 & 0.0527 \\ 0.0010 & 0.1343 & 0.0630 & 0.9322 \end{pmatrix}$$

The eigenvector corresponding to the eigenvalue  $\lambda = 1$ :

$$\pi = (0.1705, 0.0560, 0.3421, 0.4314)^{T}$$

- 1. Describe the squirrel population in the regions over the long run.
- 2. Determine and interpret the mean recurrence times

### Mean First Passage Time

Method to calculate mean first passage time and time until absorption:

$$M = (\mu_{ij}) = \begin{pmatrix} \mu_{11} & \mu_{12} & \cdots & \mu_{1N} \\ \mu_{21} & \mu_{22} & \cdots & \mu_{2N} \\ \vdots & \vdots & \cdots & \vdots \\ \mu_{N1} & \mu_{N2} & \cdots & \mu_{NN} \end{pmatrix}$$

M is the matrix of mean first passage times. The time it takes to go from  $i \rightarrow j$  is described below:

$$\mu_{ji} = p_{ji} + \sum_{k=1, k \neq j}^{N} p_{ki} (1 + \mu_{jk}) = 1 + \sum_{k=1, k \neq j}^{N} p_{ki} \mu_{jk}$$

where j is reached in 1 time step with probability  $p_{ji}$  or it takes multiple time steps and goes through state k. In Matrix form (E is matrix of 1s):

$$M = E + (M - \operatorname{diag}(M))P$$

Can solve this system ( $N^2$  equations and  $N^2$  unknowns)

### Mean First Passage Time

Suppose the MC has k absorbing states. Partition the matrix into k absorbing states and m - k transient states:

$$P = \begin{pmatrix} I & A \\ 0 & T \end{pmatrix}$$

#### Lemma 2.2

Submatrix  $T = (t_{jk})$  of transition matrix P, where indices j, k are from the set of transient states has the following property

$$\lim_{n\to\infty} T^n = \mathbf{0}.$$

### Mean First Passage Time

Let  $v_{ij}$  be the random variable for the # of visits (before absorption) to the transient state *i* beginning from *j*. The expected # of visits to *i* from *j* is

$$(E[v_{ij}]) = I + T + T^2 + T^3 + \dots = (I - T)^{-1}$$

**Fundamental Matrix** 

This is the Fundamental Matrix in DTMC:

$$F = (I - T)^{-1}$$

The expected time to absorption is the time spent in each of the transient states. Therefore, we can calculate the time to absorption by summing the columns of F:

### Expected Time Untill Absorptions

$$m = \mathbf{1}^T F$$

where 1 is a column vector of ones.

# Example

Consider the MC with transition matrix:

$$P = \begin{pmatrix} 1/2 & 0 & 0 & 1/2 \\ 1/2 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1/2 \\ 0 & 0 & 1 & 0 \end{pmatrix}$$

- 1. What are the communication classes? Classify each class?
- 2. Reorder the states and determine matrix T.
- 3. What is the fundamental matrix?
- 4. What is the mean time until absorption?

- An allele is a variant form of a gene
- Suppose there are 2 types of alleles for a given gene: a and A
- ► A diploid individual (2 sets of chromosomes) can have 3 different genotypes or combinations of alleles: *AA*, *aa*, *Aa*
- Assume 2 individuals are randomly mated. Then the next generation of their offsprings (brother and sister) are randomly mated. This inbreeding process continues each year.

- Let the mating types be states of a DTMC
- ► There are 6 states:
  - 1.  $AA \times AA$ 4.  $Aa \times aa$ 2.  $AA \times Aa$ 5.  $AA \times aa$ 3.  $Aa \times Aa$ 6.  $aa \times aa$



- Suppose parents are type 1:  $AA \times AA$
- Next generation of offsprings are all AA
- Next generation mating combinations are all type 1

▶ 
$$p_{11} = 1$$



- Suppose parents are type 2: AA × Aa
- Next generation of offsprings are 1/2 AA and 1/2 Aa
- Next generation mating combinations are (AA × AA), (AA × Aa), (Aa × Aa)
  - proportion of matings of type  $(AA \times AA) = (\frac{1}{2})(\frac{1}{2}) = \frac{1}{4}$
  - proportion of matings of type  $(AA \times Aa)^* = (\frac{1}{2})(\frac{1}{2}) + (\frac{1}{2})(\frac{1}{2}) = \frac{1}{2}$
  - proportion of matings of type  $(Aa \times Aa) = (\frac{1}{2})(\frac{1}{2}) = \frac{1}{4}$

• 
$$p_{12} = 1/4, p_{22} = 1/2, p_{32} = 1/4$$

 $^{*}(AA \times Aa)$  and  $(Aa \times AA)$ 

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- Suppose parents are type 3: Aa × Aa
- ▶ Next generation of offsprings are 1/4 AA, 1/2 Aa, and 1/4 aa
- Next generation mating combinations are (AA × AA), and (aa × aa).
  - proportion of matings of type  $(AA \times AA) = (\frac{1}{4})(\frac{1}{4}) = \frac{1}{16}$
  - proportion of matings of type  $(AA \times Aa) = \left(\frac{1}{4}\right) \left(\frac{1}{2}\right) + \left(\frac{1}{2}\right) \left(\frac{1}{4}\right) = \frac{1}{4}$
  - proportion of matings of type  $(Aa \times Aa) = (\frac{1}{2})(\frac{1}{2}) = \frac{1}{4}$
  - proportion of matings of type  $(Aa \times aa) = \left(\frac{1}{2}\right) \left(\frac{1}{4}\right) + \left(\frac{1}{4}\right) \left(\frac{1}{2}\right) = \frac{1}{4}$
  - proportion of matings of type  $(AA \times aa) = \left(\frac{1}{4}\right) \left(\frac{1}{4}\right) + \left(\frac{1}{4}\right) \left(\frac{1}{4}\right) = \frac{1}{8}$
  - proportion of matings of type  $(aa \times aa) = (\frac{1}{4})(\frac{1}{4}) = \frac{1}{16}$

▶ 
$$p_{13} = 1/16, p_{23} = 1/4, p_{33} = 1/4, p_{43} = 1/4, p_{53} = 1/8, p_{63} = 1/16$$

$$p = \begin{pmatrix} 1 & 1/4 & 1/16 & 0 & 0 & 0 \\ 0 & 1/2 & 1/4 & 0 & 0 & 0 \\ 0 & 1/4 & 1/4 & 1/4 & 1 & 0 \\ 0 & 0 & 1/4 & 1/2 & 0 & 0 \\ 0 & 0 & 1/8 & 0 & 0 & 0 \\ 0 & 0 & 1/16 & 1/4 & 0 & 1 \end{pmatrix}$$
$$p = \begin{pmatrix} 1 & | & 1/4 & 1/16 & 0 & 0 & | & 0 \\ - & - & - & - & - & - & - \\ 0 & | & 1/2 & 1/4 & 0 & 0 & | & 0 \\ 0 & | & 1/4 & 1/4 & 1/4 & 1 & | & 0 \\ 0 & | & 0 & 1/4 & 1/2 & 0 & | & 0 \\ 0 & | & 0 & 1/8 & 0 & 0 & | & 0 \\ - & - & - & - & - & - & - \\ 0 & | & 0 & 1/16 & 1/4 & 0 & | & 1 \end{pmatrix} = \begin{pmatrix} 1 & A & 0 \\ 0 & T & 0 \\ 0 & B & 0 \end{pmatrix}$$

### Questions

- 1. What are the communication classes? Classify them.
- 2. Are there any absorbing states?
- 3. Determine the expected time until absorption.
- 4. What are the probabilities of absorption into states 1 and 6?

# Unrestricted Random Walk in Higer Dimensions

Recall infinite random walk in 1D:



p > 0 is the probability of moving to the right:  $p_{i+1,i} = p$ q > 0 is the probability of moving to the left:  $p_{i,i+1} = q$ p + q = 1

### Properties

- 1. MC is irreducible
- 2. MC is periodic with period =2
- 3. Assume  $p = q = \frac{1}{2}$ , then the MC is recurrent
- 4. Assume  $p \neq q$ , then the MC is transient

# Unrestricted Random Walk in Higer Dimensions

- 1 Dimension:
  - Chain is null recurrent if and only if p = q = 1/2
  - Probability of moving left equals probability of moving right
- 2 Dimensions:
  - If probabilities of moving in any direction are equal (1/4 for up, down, right, and left) then the chain is null recurrent
- 3 Dimensions:
  - If probabilities of moving in any direction are equal (1/6 for up, down, right, left, forward, and backward) then the chain is transient
  - A path along a line or in a plane is much more restricted than in space
  - Behavior is 3 or higher dimensions in more complicated